#### Modeling for Discrete Features Hidden Markov Models I

Matthew Stone CS 520, Spring 2000 Lecture 8

# Relaxing Independence Assumptions

Want to specify

$$P(\mathbf{x} \mid \omega_i)$$
  $\mathbf{x} \in \Delta^k$ 

- few parameters for training and inference
- but accurate representation of distribution
- Seen two extremes
  - full list and naïve Bayes

# Relaxing Independence Assumptions

- Intermediate specs depend on problem
- Start with an important special case: sequential features
- Key assumption: Markov property
  - At each step in the sequence, the state depends only on the previous state

#### Some Terminology

- We'll reserve class or category to refer to the c alternative ω<sub>i</sub>
- We'll use state to refer to the changing variable that governs successive features
  - concrete possible states:  $\delta_1, \delta_2, \cdots$
  - event of being in state i at step t:  $\delta_i^{(t)}$
  - variable for events at step t.  $\delta^{(t)}$
  - variable over sequences of events:  $\delta$

### Simple Question

Say we observe a state sequence directly

$$\boldsymbol{x} = \boldsymbol{\delta} = \left\langle \delta^{(1)}, \delta^{(2)}, \cdots, \delta^{(m)} \right\rangle$$

Must model how likely x is for this class

$$P(\delta \mid \omega_i, len = m)$$

(We restrict attention to sequences of length *m* for ease of normalization.)

For Bayes discrimination

$$P(\omega_i \mid \delta) \propto P(\delta \mid \omega_i, \text{len} = m)P(\omega_i)$$

#### Modeling

• Factor in the causal direction:

$$P(\delta) = P(\delta^{(1)}) \prod_{t=2}^{m} P(\delta^{(t)} | \delta^{(1)}, \dots, \delta^{(t-1)})$$

• Markov property, I:  $\delta^{(t)}$  depends only on  $\delta^{(t-1)}$ 

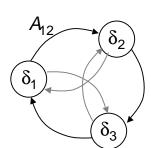
$$P(\delta) = P(\delta^{(1)}) \prod_{t=2}^{m} P(\delta^{(t)} | \delta^{(t-1)})$$

• Markov property, II:

$$P(\delta^{(t)} | \delta^{(t-1)})$$
 does not vary with  $t$ 

#### Visualization

• Diagram of states and arcs



Arcs determine matrix A

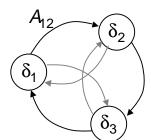
$$A_{ij} = P(\delta_j^{(t)} \mid \delta_i^{(t-1)})$$

Meas. x gives events, e.g.:

$$\boldsymbol{x} = \left\langle \delta_3^{\,(1)}, \delta_1^{\,(2)}, \delta_2^{\,(3)}, \delta_1^{\,(4)} \right\rangle$$

#### Visualization

• Diagram of states and arcs



Arcs determine matrix A

$$A_{ij} = P(\delta_j^{(t)} \mid \delta_i^{(t-1)})$$

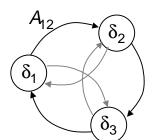
x determines arcs used

$$A^{[\mathbf{x},t]} := A_{ij} \text{ such that}$$

$$\left\langle x^{(t-1)}, x^{(t)} \right\rangle = \left\langle \delta_i, \delta_j \right\rangle$$

#### Visualization

• Diagram of states and arcs

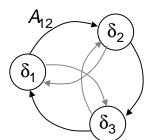


Calculate P(x) by

$$P(\mathbf{x}) = P(x^{(1)}) \prod_{t=2}^{m} P(x^{(t)} \mid x^{(t-1)})$$
$$= P(x^{(1)}) \prod_{t=2}^{m} A^{[\mathbf{x},t]}$$

#### Visualization

• Diagram of states and arcs



Example: x gives path

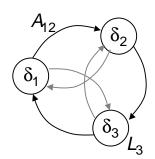
$$\boldsymbol{x} = \left\langle \delta_3^{(1)}, \delta_1^{(2)}, \delta_2^{(3)}, \delta_1^{(4)} \right\rangle$$

$$P(\mathbf{x}) = P(x^{(1)}) \prod_{t=2}^{m} A^{[\mathbf{x},t]}$$

$$= P(\delta_3^{(1)}) A_{31} A_{12} A_{21}$$

#### Visualization

• Diagram of states and arcs



Info about initial state

$$L_i = P(\delta_i^{(1)})$$

Example, ctd:

$$P(\mathbf{x}) = P(\delta_3^{(1)}) A_{31} A_{12} A_{21}$$
$$= L_3 A_{31} A_{12} A_{21}$$

**Notation**  $L^{[x]} := L_i$  if  $x^{(1)} = \delta_i$ 

#### Classification Situation

• Distribution on measurements by class

$$P(\mathbf{x} \mid \omega_i, \text{len} = m)$$

- Given by
  - Priors on initial states L
  - Transition matrix A
- Assuming
  - Set of n (observable) states  $\Delta$
  - Fixed length *m* for sequences

# Classification Situation (CONTINUED)

- Opportunities for finer representation
  - Naïve Bayes has *m*(*n*-1) parameters
  - Markov model has *n*(*n*+1) parameters
- Better independence assumptions

#### Markov Model Uses

- There are some problems where you can measure the changing state directly
  - text compression
  - correcting text (OCR, typographical errors)

# Markov Model Uses (CONTINUED)

- Treat texts as word sequences
  - set  $\Delta$  of observations (and states): words
  - matrix A contains estimates of
     bigram frequency by class probability,
     given you see word i now, of seeing word j
     immediately following
  - obtained from training sequences in class by counting and smoothing

#### But

- In the more frequent case:
  - You can't observe the state directly -
  - You must infer the state given indirect measurements
- Hidden Markov Models (HMMs) take this into account

### **Extended Terminology**

- We retain states and state variables:
  - event of being in state *i* at step t:  $\delta_i^{(t)}$
  - variable for events at step t.  $\delta^{(t)}$
- We observe a symbol at each step:
  - concrete symbols:  $v_1, v_2, \cdots$
  - event of observing i at step t:  $v_i^{(t)}$
  - variable for symbol at step t:  $v^{(t)}$
  - variable over observed sequences: v

#### **Extended Assumption**

- The symbol observed at time t depends only on the state at time t
  - and does not vary with t
  - specified by matrix B

$$B_{jk} = P(v_k^{(t)} | \delta_j^{(t)})$$

$$B^{[\mathbf{v}, \delta, t]} := B_{jk} \text{ such that }$$

$$\left\langle \delta^{(t)}, v^{(t)} \right\rangle = \left\langle \delta_j, v_k \right\rangle$$

#### **HMM Trajectory**

- Three problems must be solved to use these more flexible models
  - Evaluation:

Compute 
$$P(\mathbf{v} \mid \omega_i, \text{len} = m)$$

– Decoding:

Find argmax 
$$P(\delta | \mathbf{v}, \omega_i)$$

- Learning:

Train A and B given observations only

### Evaluation – Theory

• List all s state sequences with m elements:

$$\mathbf{s}_1, \mathbf{s}_2, \cdots, \mathbf{s}_s$$

• Use Markov assumption to find:

$$P(\mathbf{s}_a) = L^{[\mathbf{s}_a]} \prod_{u=2}^m A^{[\mathbf{s}_a,u]}$$

• Use observation assumption to find:

$$P(\mathbf{v} \mid \mathbf{s}_a) = \prod_{u=1}^m B^{[v, \mathbf{s}_a, u]}$$

# Evaluation – Theory (CONTINUED)

Given observation v:

$$P(\mathbf{v}, \mathbf{s}_u) = P(\mathbf{v} \mid \mathbf{s}_u) P(\mathbf{s}_u)$$
$$= L^{[\mathbf{s}_a]} \prod_{u=2}^m A^{[\mathbf{s}_a, u]} \prod_{u=1}^m B^{[v, \mathbf{s}_a, u]}$$

• Thus -

$$P(\mathbf{v}) = \sum_{a=1}^{s} \left( L^{[\mathbf{s}_a]} \prod_{u=2}^{m} A^{[\mathbf{s}_a, u]} \prod_{u=1}^{m} B^{[v, \mathbf{s}_a, u]} \right)$$

# Fortunately We Can Table Sums

- First, two pieces of notation
  - Probability of being in state j at step t having seen first t observations:

$$P(\delta_j^{(t)}, \mathbf{v}^{(\leq t)})$$

- Access from B:

$$B_j^{[\mathbf{v},t]} := B_{jk} \text{ if } v^{(t)} = v_k$$

– Fixing a sequence to match  $\alpha$  after t

$$\mathbf{s}^{(>t)} = \alpha$$

### **Tabling Sums**

• We find  $P(\mathbf{v}^{(\leq t)})$  as before, making an arbitrary selection among sequences:

$$P(\mathbf{v}^{(\leq t)}) = \sum_{\mathbf{s}_{a}^{(>t)} = \delta} \left( L^{[\mathbf{s}_{a}]} \prod_{u=2}^{t} A^{[\mathbf{s}_{a},u]} \prod_{u=1}^{t} B^{[v,\mathbf{s}_{a},u]} \right)$$

· Narrow to one state by restricting sum:

$$P(\delta_j^{(t)}, \mathbf{v}^{(\leq t)}) = \sum_{\mathbf{s}_a^{(>t-1)} = \delta'} \left( L^{[\mathbf{s}_a]} \prod_{u=2}^t A^{[\mathbf{s}_a, u]} \prod_{u=1}^t B^{[v, \mathbf{s}_a, u]} \right)$$

• Ensure match with  $\delta' := \delta[t : \delta_i]$ 

# Tabling Sums (CONTINUED)

Take current formula:

$$P(\delta_{j}^{(t)}, \mathbf{v}^{(\leq t)}) = \sum_{\mathbf{s}_{a}^{(>t-1)} = \delta'} \left( L^{[\mathbf{s}_{a}]} \prod_{u=2}^{t} A^{[\mathbf{s}_{a}, u]} \prod_{u=1}^{t} B^{[v, \mathbf{s}_{a}, u]} \right)$$

• And condition on t-1:  $s_a^{(>t-2)} = \delta[t-1:\delta_i]$ 

$$=\sum_{i=1}^{n}\left[\sum_{\mathbf{s}_{a}^{(>t-2)}=\delta'[t-1:\delta_{i}]}\left(L^{\left[\mathbf{s}_{a}\right]}\prod_{u=2}^{t}A^{\left[\mathbf{s}_{a},u\right]}\prod_{u=1}^{t}B^{\left[v,\mathbf{s}_{a},u\right]}\right)\right]$$

# Tabling Sums (CONTINUED)

• Rewrite:

$$= \sum_{i=1}^{n} \left[ \sum_{\mathbf{s}_{a}^{(>t-2)} = \delta'[t-1:\delta_{i}]} \left( L^{[\mathbf{s}_{a}]} A^{[\mathbf{s}_{a},t]} B^{[\mathbf{v},\mathbf{s}_{a},t]} \prod_{u=2}^{t-1} A^{[\mathbf{s}_{a},u]} \prod_{u=1}^{t-1} B^{[\mathbf{v},\mathbf{s}_{a},u]} \right) \right]$$

And factor:

$$= \sum_{i=1}^{n} \left[ A_{ij} B_{j}^{[\mathbf{v},t]} \sum_{\mathbf{s}_{a}^{(>t-2)} = \delta'[t-1:\delta_{j}]} \left( L^{[\mathbf{s}_{a}]} \prod_{u=2}^{t-1} A^{[\mathbf{s}_{a},u]} \prod_{u=1}^{t-1} B^{[\mathbf{v},\mathbf{s}_{a},u]} \right) \right]$$

# Tabling Sums (CONTINUED)

· And factor again:

$$P(\delta_j^{(t)}, \mathbf{v}^{(\leq t)}) = \sum_{j=1}^n \left[ A_{ij} B_j^{[\mathbf{v}, t]} P(\delta_i^{(t-1)}, \mathbf{v}^{(\leq t-1)}) \right]$$

### The Big Picture

• Recurrence says how to step forward...





 $v^{(t)}$ 

Suppose you've reached step t, and you've tabulated the probabilities of seeing each of the possible states...

## The Big Picture

• Recurrence says how to step forward...

$$\delta_1 \qquad p_1 = P(\delta_1^{(t)}, \mathbf{v}^{(\leq t)})$$

••• 
$$(\delta_2)$$
  $p_2 = P(\delta_2^{(t)}, \mathbf{v}^{(\leq t)})$ 

$$\begin{array}{cc}
\delta_2 & \rho_2 = P(\delta_2^{(t)}, \mathbf{v}^{(\leq t)}) \\
\hline
\delta_3 & \rho_3 = P(\delta_3^{(t)}, \mathbf{v}^{(\leq t)})
\end{array}$$

 $v^{(t)}$ 

...like so

## The Big Picture

- Recurrence says how to step forward...
  - $p_1(\delta_1)$
- $\delta_1$
- $p_2(\delta_2)$
- $\left(\delta_{2}\right)$
- $p_3(\delta_3)$
- $\delta_3$
- $v^{(t)}$
- $v^{(t+1)}$

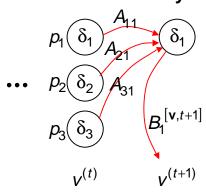
Consider the output symbol at the next step, and each of the states that might have produced it.

## The Big Picture

- Recurrence says how to step forward...
  - $p_1(\delta_1)$
- $\delta_1$  ? =  $P(\delta_1^{(t+1)}, \mathbf{v}^{(\leq t+1)}) = p_1'$
- $p_2(\delta_2)$
- $\delta_2$  ? =  $P(\delta_2^{(t+1)}, \mathbf{v}^{(\leq t+1)}) = p_2'$
- $p_3(\delta_3)$
- $\delta_3$  ? =  $P(\delta_3^{(t+1)}, \mathbf{v}^{(\leq t+1)}) = p_3'$
- $v^{(t)}$
- $v^{(t+1)}$  Want to assign probabilities to the new states.

### The Big Picture

• Recurrence says how to step forward...



For each new state

$$p'_{j} = \sum_{i=1}^{n} \left[ A_{ij} B_{j}^{[\mathbf{v},t+1]} p_{i} \right]$$

$$p'_{j} = \sum_{i=1}^{n} \left[ A_{ij} B_{j}^{[\mathbf{v},t+1]} p_{i} \right]$$
E.g.:
$$p'_{1} = A_{11} B_{1}^{[\mathbf{v},t+1]} p_{1} + A_{12} B_{1}^{[\mathbf{v},t+1]} p_{2} + A_{13} B_{1}^{[\mathbf{v},t+1]} p_{3}$$

### **Evaluation – Summary**

- · We have defined and justified
  - HMM forward algorithm
  - determining probabilities of observations
- Build table
  - Initialize:  $p_{j,0} = L_j B_j^{[\mathbf{v},0]}$
  - Step forward:  $p_{j,t+1} = \sum_{i=1}^{n} [A_{ij}B_{j}^{[\mathbf{v},t+1]}p_{i,t}]$
  - Finish:  $P(\mathbf{v} \mid \text{len} = m) = \sum_{i=1}^{n} p_{i,m}$

#### **Use of Evaluation**

• We have c models

$$\omega_a \Rightarrow \left\langle \mathbf{L}^a, \mathbf{A}^a, \mathbf{B}^a \right\rangle$$

- Each model represents distribution over sequences in the class, e.g. –
  - likely word sequences
  - likely sound sequences for saying a word
  - likely motion patterns in gesture

# Use of Evaluation (CONTINUED)

- We get some observed sequence v
- We can classify v by Bayes's formula:

Choose  $\underset{\omega_a}{\operatorname{argmax}} P(\mathbf{v} \mid \omega_a) P(\omega_a)$  where  $P(\mathbf{v} \mid \omega_a)$  is got by HMM forward