## Bayesian Decision Theory THE GENERAL CASE

- Finite set of c states of nature
  - $-\omega_1...\omega_c$
  - priors  $P(\omega_1)$ ...  $P(\omega_c)$
- · Measurement is a feature vector
  - $-\mathbf{x} \in \Re^{k}$
  - k-dimensional Euclidean feature space
  - likelihoods  $p(\mathbf{x}|\omega_1)... p(\mathbf{x}|\omega_c)$

## Bayesian Decision Theory THE GENERAL CASE

- a different actions are possible
  - $-\alpha_1...\alpha_a$
  - Loss functions  $\lambda(\alpha_1|\omega_1)...\lambda(\alpha_a|\omega_c)$

## Bayesian Decision Theory THE GENERAL CASE

• Bayes's formula again gives:

$$P(\omega_j \mid \mathbf{x}) = \frac{p(\mathbf{x} \mid \omega_j)P(\omega_j)}{p(\mathbf{x})}$$

• Evidence now is:

$$p(\mathbf{x}) = \sum_{j=1}^{c} p(\mathbf{x} \mid \omega_{j}) P(\omega_{j})$$

### Bayesian Decision Theory THE GENERAL CASE

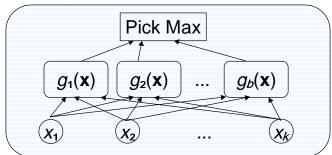
• Risk (expected loss) defined as:

$$R(\alpha_i \mid \mathbf{x}) = \sum_{j=1}^{c} \lambda(\alpha_i \mid \omega_j) P(\omega_j \mid \mathbf{x})$$

- Decision algorithm:
  - given x
  - choose  $\alpha_i$  for which  $R(\alpha_i | \mathbf{x})$  is minimum

#### **Building classifiers**

 Bayesian decision theory leads to the following picture of a classifier:



• g<sub>i</sub> are called discriminant functions

### Sample Discriminant Functions

• Risk:

$$g_{i}(\mathbf{x}) = -R(\alpha_{i}|\mathbf{x})$$
Pick Max
$$g_{1}(\mathbf{x}) \quad g_{2}(\mathbf{x}) \quad \dots \quad g_{a}(\mathbf{x})$$

$$x_{1} \quad x_{2} \quad \dots \quad x_{k}$$

### Sample Discriminant Functions

• Likelihood:

$$g_{i}(\mathbf{x}) = P(\omega_{i}|\mathbf{x})$$
Pick Max
$$g_{1}(\mathbf{x}) \quad g_{2}(\mathbf{x}) \quad \dots \quad g_{c}(\mathbf{x})$$

$$x_{1} \quad x_{2} \quad \dots \quad x_{k}$$

### Sample Discriminant Functions

• Non-normalized likelihood:

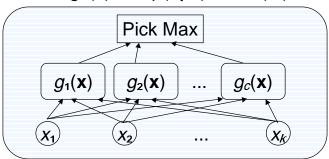
$$g_{i}(\mathbf{x}) = p(\mathbf{x}|\omega_{i})P(\omega_{i})$$
Pick Max
$$g_{1}(\mathbf{x}) \quad g_{2}(\mathbf{x}) \quad \dots \quad g_{c}(\mathbf{x})$$

$$\chi_{2} \quad \dots \quad \chi_{k}$$

#### Sample Discriminant Functions

• Non-normalized log likelihood:

$$g_i(\mathbf{x}) = \ln p(\mathbf{x}|\omega_i) + \ln P(\omega_i)$$



# Implementing discriminant functions

- Encodes and exploits assumptions about the distributions of measurements
- Important special case: normally distributed measurements

## Understanding distributions CONTINUOUS SCALAR CASE

• Expected value of a scalar function

$$E[f(x)] \equiv \int f(x) p(x) dx$$

Mean (expected value of x)

$$\mu = E[x] = \int x \, p(x) \, dx$$

• Variance (expected squared deviation)

$$\sigma^2 = E[(x - \mu)^2] = \int (x - \mu)^2 p(x) dx$$

Entropy (negative expected log density)

$$H(p(x)) = -E[\ln p(x)] = -\int p(x) \ln p(x) dx$$

#### **Univariate Normal Density**

Defined as

$$p(x) = \frac{1}{\sqrt{2\pi\sigma}} \exp \left[ -\frac{1}{2} \left( \frac{x - \mu}{\sigma} \right)^2 \right]$$

· Shorthand:

$$p(x) \sim N(\mu, \sigma^2)$$

### Simple Classification Example

- Binary decision from one measurement
  - $-P(\omega_1), P(\omega_2)$
  - $-p(x|\omega_1) \sim N(\mu_1,\sigma_1^2), p(x|\omega_2) \sim N(\mu_2,\sigma_2^2)$
  - $-g_i(x) = \ln p(x|\omega_i) + \ln P(\omega_i)$

#### Working out the details...

• Calculate  $g_i(x)$  as:

$$g_{i}(x) = \ln \left( \frac{1}{\sqrt{2\pi}\sigma_{i}} \exp \left[ -\frac{1}{2} \left( \frac{x - \mu_{i}}{\sigma_{i}} \right)^{2} \right] \right) + \ln P(\omega_{i})$$

$$= -\frac{1}{2} \ln 2\pi - \ln \sigma_{i} - \frac{1}{2} \left( \frac{x - \mu_{i}}{\sigma_{i}} \right)^{2} + \ln P(\omega_{i})$$

$$\approx -\frac{(x - \mu_{i})^{2}}{2\sigma_{i}^{2}} - \ln \sigma_{i} + \ln P(\omega_{i})$$

#### Decide $\omega_1$ if $g_1(x) > g_2(x)$

• In other words if:

$$-\frac{(x-\mu_{1})^{2}}{2\sigma_{1}^{2}} - \ln\left(\frac{\sigma_{1}}{P(\omega_{1})}\right) > -\frac{(x-\mu_{2})^{2}}{2\sigma_{2}^{2}} - \ln\left(\frac{\sigma_{2}}{P(\omega_{2})}\right)$$

$$\frac{(x-\mu_{1})^{2}}{2\sigma_{1}^{2}} < \frac{(x-\mu_{2})^{2}}{2\sigma_{2}^{2}} + \ln\left(\frac{\sigma_{2}P(\omega_{1})}{\sigma_{1}P(\omega_{2})}\right)$$

$$\frac{(x-\mu_{1})^{2}}{2\sigma_{1}^{2}} < \frac{(x-\mu_{2})^{2}}{2\sigma_{2}^{2}} + t$$

## Understanding distributions MULTIVARIATE CASE

- Expected value of a scalar function
  - integrate over the whole feature space:

$$E[f(\mathbf{x})] \equiv \int f(\mathbf{x}) p(\mathbf{x}) d\mathbf{x}$$

- For vectors, matrices acts componentwise
  - Mean (expected value of x)

$$\mu = E[\mathbf{x}] = \int \mathbf{x} \ p(\mathbf{x}) \ d\mathbf{x}$$
  
$$\mu_i = E[x_i] = \int x_i \ p(\mathbf{x}) \ d\mathbf{x}$$

## Understanding distributions MULTIVARIATE CASE, CTD

• Covariance (deviation+correlation):

$$\Sigma = E[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^{\mathsf{T}}] = \int (\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^{\mathsf{T}} \rho(\mathbf{x}) d\mathbf{x}$$
  
$$\sigma_{ij} = E[(x_i - \boldsymbol{\mu}_i)(x_j - \boldsymbol{\mu}_j)] = \int (x_i - \boldsymbol{\mu}_i)(x_j - \boldsymbol{\mu}_j) \rho(\mathbf{x}) d\mathbf{x}$$

#### **Multivariate Normal Density**

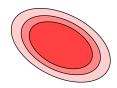
Defined as

$$p(\mathbf{x}) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp\left[-\frac{1}{2} (\mathbf{x} - \mu)^{\mathsf{T}} \Sigma^{-1} (\mathbf{x} - \mu)\right]$$

• Shorthand:

$$p(\mathbf{x}) \sim N(\mu, \Sigma)$$

#### A Quick Visualization



 Probability density falls off in hyperellipsoids of constant Mahalonobis distance

$$r^2 = (\mathbf{x} - \mathbf{\mu})^{\mathsf{T}} \Sigma^{-1} (\mathbf{x} - \mathbf{\mu})$$

• Covariance determines rotation and shape

# Sample multivariate classification case

- Features are statistically independent, across categories
- Features have the same variance  $\sigma^2$





• Geometric intuition: categories determine equal-size hyperspherical clusters

#### Formal description

· Conditional pdf is normal for each class

$$p(\mathbf{x} \mid \omega_i) \sim N(\mu_i, \Sigma)$$

- mean vector for each class: µ/
- covariance matrix  $\Sigma = \sigma^2$
- determinant  $|\Sigma| = \sigma^{2k}$
- inverse  $\Sigma^{-1} = (1/\sigma^2)I$

#### Working it through

• By algebra as in univariate case we get:

$$g_i(\mathbf{x}) = -\frac{\|\mathbf{x} - \boldsymbol{\mu}\|^2}{2\sigma^2} + \ln P(\omega_i)$$

But it's not necessary to compute distances:

$$g_i(\mathbf{x}) = -\frac{1}{2\sigma^2} \left[ \mathbf{x}^\mathsf{T} \mathbf{x} - 2\mu^\mathsf{T} \mathbf{x} + \mu^\mathsf{T} \mu \right] + \ln P(\omega_i)$$

 $\mathbf{x}^{\mathsf{T}}\mathbf{x}$  is the same for all *i* 

#### **Linear Discriminant Functions**

· So equivalently:

$$g_i(\mathbf{x}) = \mathbf{w}_i^{\mathsf{T}} \mathbf{x} + w_{i0}$$

• For weight vector

$$\mathbf{w}_{i}^{\mathsf{T}} = \frac{1}{\sigma^{2}} \mu_{i}$$

· and threshold or bias

$$W_{i0} = \frac{-1}{2\sigma^2} \mu^T \mu + \ln P(\omega_i)$$

#### The behavior of classifiers

- Decision rule divides feature space into decision regions
  - If  $g_i(\mathbf{x}) > g_j(\mathbf{x})$  for all j, then  $\mathbf{x}$  is in region i
  - Regions separated by decision boundaries (where largest discriminant functions tie)

#### The behavior of linear classifiers

- Decision surfaces are pieces of hyperplanes  $g_i(\mathbf{x}) = g_j(\mathbf{x})$ 
  - Orthogonal to the line between the means
  - Shifted from halfway by variance and priors
- Explicitly:  $\mathbf{w}^{\mathsf{T}}(\mathbf{x} \mathbf{x}_0) = 0$

$$\mathbf{w} = \mu_i - \mu_j$$

$$\mathbf{x}_{0} = \frac{1}{2} \left( \mu_{i} + \mu_{j} \right) - \frac{\sigma^{2}}{\left\| \mu_{i} - \mu_{j} \right\|^{2}} \ln \frac{P(\omega_{i})}{P(\omega_{j})} \left( \mu_{i} - \mu_{j} \right)$$